

Evolution of cooperation in finite populations

Sabin Lessard
Université de Montréal

Contents

1. Examples of cooperation
2. Prisoner's Dilemma
3. Evolutionary dynamics in an infinite population
4. Fixation probability in a finite population
5. One-third law of evolution
6. Projected average excess in payoff
7. Group-structured population
8. Differential contributions of groups
9. Skewed contributions of groups
10. Summary and comments

Thanks

Collaborators

Véronique Ladret
Philippe Lahaie
Samuel Langevin
David Lasalle-Ialongo
Géraldine Martin

Competitors

Martin Nowak
François Rousset
and many others



1. Examples of cooperation in nature













Cooperation is widespread...

but emergence of cooperation is problematic!

- ▶ A **game-theoretic framework** based on pairwise interactions is a first step toward a better understanding of a complex phenomenon.
- ▶ This framework can provide **clues about favorable conditions** for the evolution of cooperation.

2. Prisoner's Dilemma (PD)



Payoff Matrix

<i>Cooperate</i>	<i>Reward</i>	<i>Sucker's payoff</i>
<i>Defect</i>	<i>Temptation</i>	<i>Punishment</i>
against	<i>Cooperate</i>	<i>Defect</i>

$$T > R > P > S$$

C	$R = 5$	$S = 1$
D	$T = 14$	$P = 3$
against	C	D

Iterated Prisoner's Dilemma (IPD)

- ▶ PD repeated n times between the same players with additive payoffs

Tit-for-Tat (A)	$a = Rn$	$b = S + P(n - 1)$
Always-Defect (B)	$c = T + P(n - 1)$	$d = Pn$
against	A	B

$$a > c > d > b \text{ as soon as } n > \frac{T-P}{R-P}$$

A	$a = 50$	$b = 28$
B	$c = 41$	$d = 30$
against	A	B

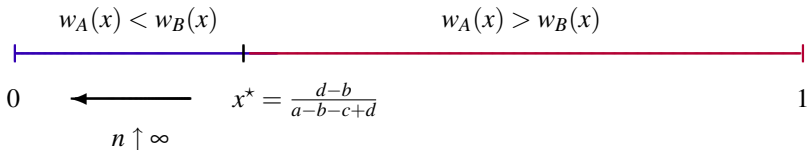
for $n = 10$ in the previous example

Expected payoffs in an infinite population

- ▶ random pairwise interactions
- ▶ x : frequency of A

$$w_A(x) = ax + b(1 - x)$$

$$w_B(x) = cx + d(1 - x)$$



3. Evolutionary dynamics in an infinite population

- ▶ discrete, non-overlapping generations
- ▶ fitness : $1 + s \times \text{payoff}$ for some intensity of selection $s \geq 0$
- ▶ $x(t)$: frequency of A in offspring in generation t before selection

$$x(t+1) = \frac{x(t)(1 + s w_A(x(t)))}{1 + s \bar{w}(x(t))}$$

$$x(t+1) - x(t) = \frac{s(a - b - c + d)x(t)(1 - x(t))(x(t) - x^*)}{1 + s \bar{w}(x(t))}$$



4. Fixation probability in a finite population

- ▶ N parents chosen at random to produce the next generation (assumed large)
- ▶ (π_1, \dots, π_N) : proportions of offspring produced in large numbers (assumed exchangeable; $\pi_i = N^{-1}$ for the Wright-Fisher model)
- ▶ $X(t)$: frequency of A in offspring in generation t before selection

In virtue of the **ergodic theorem** for Markov chains

$$X(T) \rightarrow X(\infty) = X(0) + \sum_{t \geq 0} (X(t+1) - X(t))$$

$X(\infty) = 1$ with probability $u(s)$, and 0 otherwise

► $u(s) = E_s[X(\infty)]$: probability of ultimate fixation of A

► $u(0) = X(0) = N^{-1}$

$$\begin{aligned}u(s) &= X(0) + \sum_{t \geq 0} E_s[X(t+1) - X(t)] \\&= u(0) + s(a - b - c + d) \sum_{t \geq 0} E_s \left[\frac{X(t)(1 - X(t))(X(t) - x^*)}{1 + s\bar{w}(X(t))} \right] \\&= u(0) + s(a - b - c + d) \sum_{t \geq 0} E_0[X(t)(1 - X(t))(X(t) - x^*)] + o(s)\end{aligned}$$

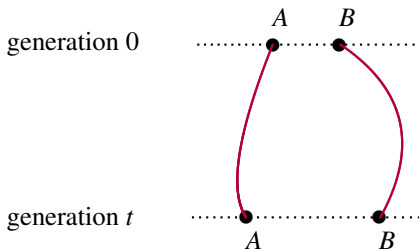
Definition: Weak selection favors A replacing B if $u(s) > u(0)$ for $s > 0$ small

$$x^* < \frac{\sum_{t \geq 0} E_0[X(t)^2(1 - X(t))]}{\sum_{t \geq 0} E_0[X(t)(1 - X(t))]} = \hat{x}$$

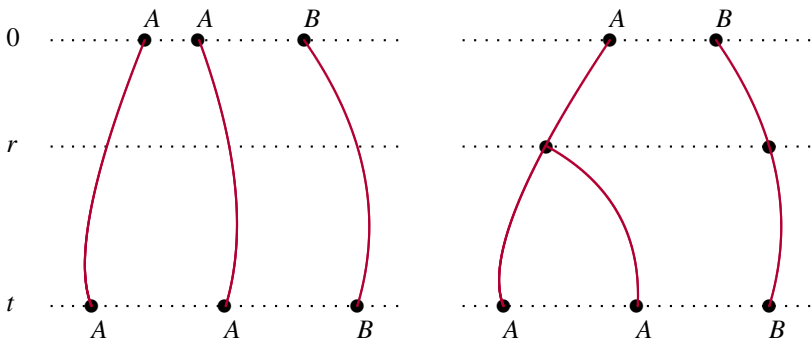
5. One-third law of evolution

$$\begin{aligned}\sum_{t \geq 0} E_0[X(t)(1 - X(t))] &= \sum_{t \geq 0} P_0(A, B \text{ in generation } t) \\ &= \sum_{t \geq 0} p_{22}(t) P_0(A, B \text{ in generation } 0) \\ &= \frac{X(0)(1 - X(0))}{1 - p_{22}}\end{aligned}$$

with $p_{22}(t) = p_{22}^t$ the probability that two offspring chosen at random in generation t descend from two distinct ancestors in generation 0



$$\begin{aligned}
\sum_{t \geq 0} E_0[X(t)^2(1 - X(t))] &= \sum_{t \geq 0} P_0(A, A, B \text{ in generation } t) \\
&= \sum_{t \geq 0} p_{33}(t) P_0(A, A, B \text{ in generation } 0) \\
&\quad + \sum_{t \geq 0} \frac{p_{32}(t)}{3} P_0(A, B \text{ in generation } 0)
\end{aligned}$$



with $p_{33}(t) = p_{33}^t$ the probability that three offspring chosen at random in generation t descend from three distinct ancestors in generation 0, and

$$p_{32}(t) = \sum_{r=0}^{t-1} p_{33}^{t-r-1} p_{32} p_{22}^r = p_{32} \left(\frac{p_{33}^t - p_{22}^t}{p_{33} - p_{22}} \right)$$

the probability that they descend from two distinct ancestors in generation 0, from which

$$\begin{aligned} \sum_{t \geq 0} E_0[X(t)^2(1 - X(t))] &= \frac{X(0)^2(1 - X(0))}{1 - p_{33}} + \frac{p_{32}X(0)(1 - X(0))}{3(1 - p_{22})(1 - p_{33})} \\ &\approx \frac{p_{32}X(0)(1 - X(0))}{3(1 - p_{22})(1 - p_{33})} \end{aligned}$$

for N large enough

For N large enough

$$\hat{x} \approx \frac{p_{32}}{3(1-p_{33})} \leq \frac{1}{3}$$

Definition (Nowak et al. *Nature* 2004): One-third law of evolution if $\hat{x} = 1/3$



Result (L. & Ladret JMB 2007): One-third law of evolution if and only if $p_{32} = 1 - p_{33}$, which means that at most 2 lineages out of 3 coalesce at a time backward in time with probability 1.

This characterizes the **Kingman coalescent** for a wide range of reproduction schemes as $N \rightarrow \infty$ with N generations as unit of time: then each pair of lineages coalesces at rate 1 independently of all others, so that

$$\lambda_{21} = 1, \lambda_{32} = 3, \lambda_{31} = 0,$$

and more generally

$$\lambda_{k,k-1} = \frac{k(k-1)}{2}$$

where λ_{kj} is the rate of death from k to j lineages backwards in time.

Skewed contributions of parents

- ▶ probability $1 - N^{-\alpha}$ that every parent produces the **same proportion** N^{-1} of offspring
- ▶ probability $N^{-\alpha}$ that one parent at random produces a **proportion ψ** of offspring and every other parent a proportion $(1 - \psi)(N - 1)^{-1}$
- ▶ in the case $\alpha < 1$ with N^α generations as unit of time as $N \rightarrow \infty$, $\lambda_{21} = \psi^2$, $\lambda_{31} = \psi^3$, $\lambda_{32} = 3\psi^2(1 - \psi)$, according to a Λ -coalescent

$$\hat{x} \approx \frac{\lambda_{32}}{3(\lambda_{31} + \lambda_{32})} = \frac{1 - \psi}{3 - 2\psi} < \frac{1}{3}$$

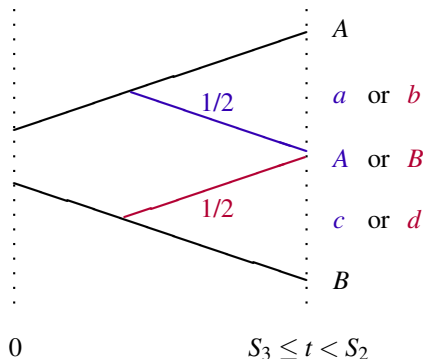
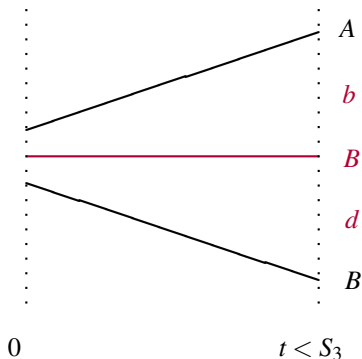
which means a **more stringent condition** for A to be favored.

6. Projected average excess in payoff

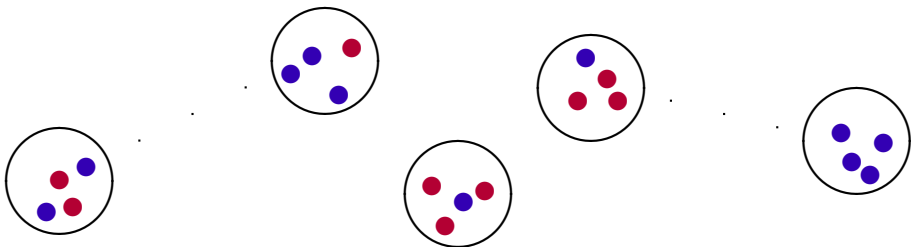
Difference between the marginal payoff to A and the mean payoff to a competitor in all generations $t \geq 0$

$$\frac{u'(0)}{X(0)} \approx (b-d)E_0(S_3) + \left[\frac{a-c}{2} + \frac{b-d}{2} \right] (E_0(S_2) - E_0(S_3))$$

with S_j for a time with j lineages, and one-third law if $E_0(S_2) = 3E_0(S_3)$



7. Group-structured population



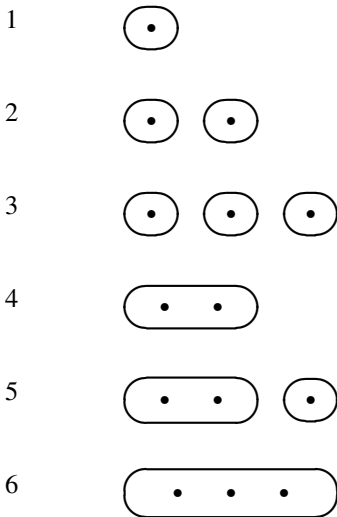
- ▶ D groups of N parents producing equal proportions of offspring (D assumed large)
- ▶ m : proportion of offspring in each group that disperse uniformly (Wright's island model)
- ▶ selection among offspring within groups after dispersal

- ▶ $X_i(t)$: frequency of A in group i before selection in generation t
- ▶ $\overline{X(t)} = D^{-1} \sum_{i=1}^D X_i(t)$: frequency of A in the population in generation t
- ▶ $\overline{X(0)} = (ND)^{-1}$

In this case weak selection favors A replacing B if $x^* < \hat{x}$

$$\begin{aligned} \hat{x} &= \frac{\sum_{t \geq 0} E_0 \left[\overline{X(t)^2(1 - X(t))} \right]}{\sum_{t \geq 0} E_0 \left[\overline{X(t)(1 - X(t))} \right]} \\ &= \frac{\sum_{t \geq 0} P_0(A, A, B \text{ in the same group in generation } t)}{\sum_{t \geq 0} P_0(A, B \text{ in the same group in generation } t)} \end{aligned}$$

States for the ancestors of 3 offspring



Transition matrix backward in time

Applying Möhle (1998) lemma

$$\mathbf{P}^{\lfloor NDf_{22}^{-1}\tau \rfloor} \rightarrow \begin{pmatrix} e^{\tau \mathbf{G}} & 0 \\ \mathbf{F}e^{\tau \mathbf{G}} & 0 \end{pmatrix}$$

as $D \rightarrow \infty$, where

$$\mathbf{F} = \begin{pmatrix} f_{21} & f_{22} & 0 \\ 0 & f_{21} & f_{22} \\ f_{31} & f_{32} & f_{33} \end{pmatrix}$$

with f_{nk} the probability for n offspring in the same group *after dispersal* to have k ancestors in different groups in the case of an infinite number of groups,

$$\mathbf{G} = \begin{pmatrix} 0 & 0 & 0 \\ 1 & -1 & 0 \\ 0 & 3 & -3 \end{pmatrix}$$

the generator for the Kingman coalescent in a well-mixed population.

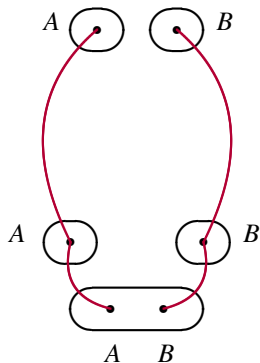
Two-time-scale argument for large D

- ▶ **Fast scattering** of lineages in the same group; **slow collecting** of lineages in different groups.
- ▶ **Times spent with lineages in the same group can be neglected** compared to times spent with lineages in different groups.
- ▶ **The expected time spent with 2 lineages in different groups** before coalescence in number of generations is approximately NDf_{22}^{-1} .

$$\sum_{t \geq 0} P_0(A, B \text{ in the same group in generation } t)$$

$$\approx \sum_{t \geq 0} P_{42}(t) P_0(A, B \text{ in different groups in generation } 0)$$

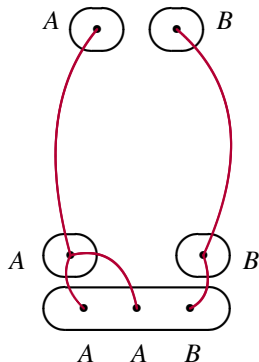
$$\approx f_{22} \times NDf_{22}^{-1} \times \overline{X(0)}(1 - \overline{X(0)})$$



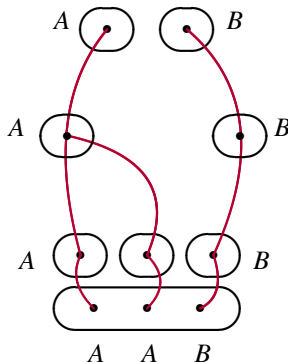
state 2 in generation 0

state 4 in generation t

$$\begin{aligned}
& \sum_{t \geq 0} P_0(A, A, B \text{ in the same group in generation } t) \\
& \approx \frac{1}{3} \sum_{t \geq 0} P_{62}(t) P_0(A, B \text{ in different groups in generation } 0) \\
& \approx \frac{1}{3} \times (f_{32} + f_{33}) \times NDf_{22}^{-1} \times \overline{X(0)}(1 - \overline{X(0)})
\end{aligned}$$



state 2 in generation 0



state 6 in generation t

$$\hat{x} \approx \frac{1-f_{31}}{3(1-f_{21})} > \frac{1}{3}$$

with

$$f_{21} = \frac{(1-m)^2}{Nm(2-m) + (1-m)^2}$$

$$f_{31} = f_{21} \left[\frac{N(1-m) + 2(N-1)(1-m)^3}{N^2m(3-3m+m^2) + (3N-2)(1-m)^3} \right]$$

which means a **less stringent condition** for A to be favored in the case of a **group-structured population**.

8. Differential contributions of groups

- Wright island model with selection among offspring before dispersal

$$\begin{aligned}u'(0) &= (b-d) \sum_{t \geq 0} E_0 \left[\overline{X(t)(1-X(t))} \right] \\&+ (a-b-c+d) \sum_{t \geq 0} E_0 \left[\overline{X(t)^2(1-X(t))} \right] \\&+ m(2-m)(b+c-2d) \sum_{t \geq 0} E_0 \left[\overline{X(t)^2} - \overline{X(t)}^2 \right] \\&+ m(2-m)(a-b-c+d) \sum_{t \geq 0} E_0 \left[\overline{X(t)^3} - \overline{X(t)} \overline{X(t)^2} \right]\end{aligned}$$

$$\begin{aligned}
E_0 \left[\overline{X(t)^2} - \overline{X(t)}^2 \right] &\approx P_0(A, B \text{ in different groups}) \\
&- P_0(A, B \text{ in the same group}) \\
E_0 \left[\overline{X(t)^3} - \overline{X(t)} \overline{X(t)^2} \right] &\approx P_0(A, A, B \text{ with } B \text{ in a different group}) \\
&- P_0(A, A, B \text{ in the same group})
\end{aligned}$$

$$\hat{x} \approx \frac{1 - \tilde{f}_{31} - m(2 - m)(\tilde{f}_{21} - \tilde{f}_{31})}{3(1 - \tilde{f}_{21})} + \frac{(a - d)(N - 1)^{-1}}{(a - b - c + d)} > \frac{1 - f_{31}}{3(1 - f_{21})}$$

where \tilde{f}_{nk} is the probability for n offspring in the same group *before dispersal* to have k ancestors in different groups with an infinite number of groups. This means an **even less stringent condition** for A to be favored.

9. Skewed contributions of groups

- ▶ selection after dispersal
- ▶ probability $D^{-\beta}$ that a group at random provides a fraction χ of migrants and every other a fraction $(1 - \chi)(D - 1)^{-1}$ for $\beta < 1$

$$u'(0) \approx \left[(b-d)f_{22} + \left(\frac{a-b-c+d}{3} \right) \left(f_{32} + f_{33} \frac{\lambda_{32}}{\lambda_{32} + \lambda_{31}} \right) \right] \\ \times \lambda_{21}^{-1} D^{\beta} \times \overline{X(0)}(1 - \overline{X(0)})$$

where λ_{kj} is the rate of transition from k to j lineages in different groups backward in time with D^{β} generations as unit of time as $D \rightarrow \infty$

$$\hat{\chi} \approx \frac{1-f_{31}-f_{33} \frac{\lambda_{31}}{\lambda_{32}+\lambda_{31}}}{3(1-f_{21})} < \frac{1-f_{31}}{3(1-f_{21})}$$

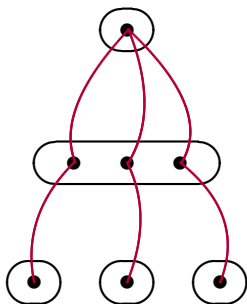
which means a **more stringent condition** for A to be favored.

Rates of transition for lineages in different groups

$$\lambda_{32} = 3(\lambda_{21} - \lambda_{31})$$

$$\lambda_{21} = (\chi m)^2 \tilde{f}_{21} = \left(\frac{\chi m}{1-m}\right)^2 f_{21}$$

$$\lambda_{31} = (\chi m)^3 \tilde{f}_{31} = \left(\frac{\chi m}{1-m}\right)^3 f_{31}$$



same ancestor in the case
of infinite number of groups

before dispersal

after dispersal

10. Summary and comments

- ▶ IPD in an **infinite population** predicts the increase in frequency of Tit-for-Tat (A) against Always-Defect (B), and therefore can explain the spread of cooperation, but only from a frequency $x > x^*$.
- ▶ IPD in a **finite population** favors a rare mutant A replacing B , and therefore can explain the advantage of cooperation to go to fixation, but only under the condition $x^* < \hat{x}$.
- ▶ In a **large population**, $\hat{x} = 1/3$ in the domain of the **Kingman coalescent**, but $\hat{x} < 1/3$ leading to a more stringent condition for cooperation to be favored if the contributions of parents in offspring are highly skewed.
- ▶ The first-order effect of selection on the probability of fixation is given by a **projected average excess in payoff**.

- ▶ In a **group-structured population** with uniform dispersal in the limit of a large number of groups, $\hat{x} > 1/3$ which means a less stringent condition for cooperation to be favored.
- ▶ The condition is weaker **if dispersal occurs after selection rather than before selection**, so that there are differential contributions of groups. On the other hand the condition is stronger **if the contributions of groups in offspring are more skewed**.
- ▶ The results obtained from the first-order effect of selection are ascertained only **under very weak selection**, actually as long as the intensity of selection is small compared to the intensity of the other evolutionary forces (drift, dispersal), but **without constraints on the reproduction scheme**.

Thanks!